## **3** Characteristic Functions

Characteristic functions (a straightforward generalization of Fourier transforms) yield a very convenient way to deal with many properties of distributions; in particular, they allow to describe weak convergence.

We use the notation  $\langle \cdot, \cdot \rangle$  and  $\|\cdot\|$  for the euclidean inner product and norm. Recall that  $\mathfrak{M}(\mathbb{R}^k)$  denotes the class of all probability measures on  $(\mathbb{R}^k, \mathfrak{B}_k)$ . Further, *i* denotes the imaginary unit. Let  $\mu \in \mathfrak{M}(\mathbb{R}^k)$ .

**Definition 1.**  $f : \mathbb{R}^k \to \mathbb{C}$  is  $\mu$ -integrable if  $\Re f$  and  $\Im f$  are  $\mu$ -integrable, in which case

$$\int f \, d\mu = \int \Re f \, d\mu + \imath \cdot \int \Im f \, d\mu.$$

A bounded continuous function  $f : \mathbb{R}^k \to \mathbb{C}$  is obviously  $\mu$ -integrable. A special class of such functions is given by  $f_y(x) := \exp(i\langle x, y \rangle)$  with  $y \in \mathbb{R}^k$ .

**Definition 2.** The mapping  $\widehat{\mu} : \mathbb{R}^k \to \mathbb{C}$  with

$$\widehat{\mu}(y) = \int \exp(i\langle x, y \rangle) \,\mu(dx), \qquad y \in \mathbb{R}^k,$$

is called the *Fourier transform* of  $\mu$ .

## Example 1.

(i) For  $\mu$  discrete,

$$\mu = \sum_{j=1}^{\infty} \alpha_j \cdot \delta_{x_j}$$

we have

$$\widehat{\mu}(y) = \sum_{j=1}^{\infty} \alpha_j \cdot \exp(i \langle x_j, y \rangle).$$

For instance, if  $\mu = Pois(\lambda)$  is the Poisson distribution with parameter  $\lambda > 0$ , then

$$\widehat{\mu}(y) = \exp(-\lambda) \cdot \sum_{j=0}^{\infty} \frac{\lambda^j}{j!} \cdot \exp(ijy) = \exp(-\lambda) \cdot \exp(\lambda \cdot \exp(iy))$$
$$= \exp(\lambda \cdot (\exp(iy) - 1)).$$

(ii) If  $\mu = f \cdot \lambda_k$  then

$$\widehat{\mu}(y) = \int \exp(i\langle x, y \rangle) \cdot f(x) \,\lambda_k(dx).$$

For any  $\lambda_k$ -integrable function f, the right-hand side defines its *Fourier trans*form, see Analysis III. For instance, if  $\mu$  is the k-dimensional standard normal distribution, i.e.,

$$f(x) = (2\pi)^{-k/2} \cdot \exp(-\|x\|^2/2),$$

then

$$\widehat{\mu}(y) = \exp(-\|y\|^2/2)$$

See Bauer (1996, p. 187) for the case k = 1. Use Fubini's Theorem for k > 1.

## Theorem 1.

- (i)  $\hat{\mu}$  is uniformly continuous on  $\mathbb{R}^k$ ,
- (ii)  $|\widehat{\mu}(y)| \leq 1 = \widehat{\mu}(0)$  for  $y \in \mathbb{R}^k$ ,
- (iii) for  $n \in \mathbb{N}$ ,  $a_1, \ldots, a_n \in \mathbb{C}$ , and  $y_1, \ldots, y_n \in \mathbb{R}^k$ ,

$$\sum_{j,\ell=1}^{n} a_j \cdot \overline{a_\ell} \cdot \widehat{\mu}(y_j - y_\ell) \ge 0$$

(positive semi-definite).

*Proof.* Ad (i): Observe that

$$\left|\exp(i\langle x, y_1\rangle) - \exp(i\langle x, y_2\rangle)\right| \le ||x|| \cdot ||y_1 - y_2||.$$

For  $\varepsilon > 0$  take r > 0 such that  $\mu(B) \ge 1 - \varepsilon$ , where  $B = \{x \in \mathbb{R}^k : ||x|| \le r\}$ . Then

$$\begin{aligned} \left| \widehat{\mu}(y_1) - \widehat{\mu}(y_2) \right| &\leq \int_B \left| \exp(i \langle x, y_1 \rangle) - \exp(i \langle x, y_2 \rangle) \right| \mu(dx) + 2 \cdot \varepsilon \\ &\leq r \cdot \|y_1 - y_2\| + 2 \cdot \varepsilon. \end{aligned}$$

Properties (ii) and (iii) are easily verified.

**Remark 1.** Bochner's Theorem states that every continuous, positive semi-definite function  $\varphi : \mathbb{R}^k \to \mathbb{C}$  with  $\varphi(0) = 1$  is the Fourier transform of a probability measure on  $(\mathbb{R}^k, \mathfrak{B}_k)$ . See Bauer (1996, p. 184) for references.

In the sequel:  $X, Y, \ldots$  are k-dimensional random vectors on a probability space  $(\Omega, \mathfrak{A}, P)$ .

**Definition 3.** The *characteristic function* of X is given by

$$\varphi_X = \widehat{P_X}.$$

Remark 2. Due to Theorem II.9.1

$$\varphi_X(y) = \int_{\mathbb{R}^k} \exp(i\langle x, y \rangle) P_X(dx) = \mathbb{E} e^{i\langle X, y \rangle}$$

Theorem 2.

(i) For every linear mapping  $T : \mathbb{R}^k \to \mathbb{R}^\ell$ 

$$\varphi_{T \circ X} = \varphi_X \circ T^t.$$

(ii) For independent random vectors X and Y

$$\varphi_{X+Y} = \varphi_X \cdot \varphi_Y.$$

In particular, for  $a \in \mathbb{R}^k$ ,

$$\varphi_{X+a} = \exp(i\langle a, \cdot \rangle) \cdot \varphi_X.$$

*Proof.* Ad (i): Let  $z \in \mathbb{R}^{\ell}$ . Use  $P_{T \circ X} = T(P_X)$  to obtain

$$\varphi_{T \circ X}(z) = \int_{\mathbb{R}^k} \exp(i\langle T(x), z \rangle) P_X(dx) = \varphi_X(T^t(z)).$$

Ad (ii): Let  $z \in \mathbb{R}^k$ . Fubini's Theorem and Theorem III.5.5 imply

$$\varphi_{X+Y}(z) = \int_{\mathbb{R}^{2k}} \exp(i\langle x+y,z\rangle) P_{(X,Y)}(d(x,y)) = \varphi_X(z) \cdot \varphi_Y(z).$$

Corollary 1 (Convolution Theorem). For probability measures  $\mu_j \in \mathfrak{M}(\mathbb{R})$ ,

$$\widehat{\mu_1 \ast \mu_2} = \widehat{\mu_1} \cdot \widehat{\mu_2}.$$

*Proof.* Use Theorem 2.(ii) and Theorem III.5.8.

**Example 2.** For  $\mu = N(m, \sigma^2)$  with  $\sigma \ge 0$  and  $m \in \mathbb{R}$ 

$$\widehat{\mu}(y) = \exp(imy) \cdot \exp(-\sigma^2 y^2/2).$$

See Example 1.(ii) and Theorem 2.

**Lemma 1.** For  $z \in \mathbb{R}$  and  $\sigma > 0$ 

$$\int \exp(-\imath y z/\sigma^2) \cdot \widehat{\mu}(y/\sigma^2) N(0,\sigma^2)(dy) = \int \exp(-(z-x)^2/(2\sigma^2)) \mu(dx).$$

*Proof.* We note first that the function  $\varphi : \mathbb{R} \to \mathbb{C}$ ,

$$\varphi(r) := \int_{\mathbb{R}} e^{-(y-ir)^2/2\sigma^2} \, dy = \int_{\Re} e^{-(y^2-r^2)/2\sigma^2 + iyr/\sigma^2} \, dy$$

is constant. (This can be seen by computing  $\varphi'$  by switching integration and differentiation, or by Cauchy's integral theorem.) The left hand side in the lemma equals

$$\int \int e^{-iyz/\sigma^2} e^{iyx/\sigma^2} \frac{1}{\sqrt{2\pi\sigma^2}} e^{-y^2/(2\sigma^2)} \, dy \, d\mu(x) = \int e^{-(z-x)^2/(2\sigma^2)} \cdot \left(\frac{1}{\sqrt{2\pi\sigma^2}} \int e^{-(y^2-(x-z)^2)/(2\sigma^2)+iy(x-z)} \, dy \, d\mu(x)\right) dy$$

The inner integral equals  $\varphi(x-z) = \varphi(0) = 1$ .

**Lemma 2.** For  $\sigma_n > 0$  with  $\lim_{n\to\infty} \sigma_n = 0$ ,

$$N(0, \sigma_n^2) * \mu \xrightarrow{\mathrm{w}} \mu.$$

*Proof.* Consider independent random variables  $X_n$  and Y such that  $X_n \sim N(0, \sigma_n^2)$ and  $Y \sim \mu$ . Then  $X_n \xrightarrow{\mathfrak{L}^2} 0$ , and therefore  $X_n + Y \xrightarrow{\mathfrak{L}^2} Y$ , which implies

$$X_n + Y \stackrel{\mathrm{d}}{\longrightarrow} Y$$

**Theorem 3 (Uniqueness Theorem).** For probability measures  $\mu_j \in \mathfrak{M}(\mathbb{R}^k)$ ,

$$\mu_1 = \mu_2 \qquad \Leftrightarrow \qquad \widehat{\mu_1} = \widehat{\mu_2}.$$

*Proof.* ' $\Rightarrow$ ' holds by definition. ' $\Leftarrow$ ': See Bauer (1996, Thm. 23.4) or Billingsley (1979, Sec. 29) for the case k > 1. Here: the case k = 1. For  $\sigma > 0$  and  $A \in \mathfrak{B}$ , Theorem III.5.9 yields

$$N(0,\sigma^2) * \mu_j(A) = (2\pi\sigma^2)^{-1/2} \cdot \int_A \int \exp(-(z-x)^2/(2\sigma^2)) \,\mu_j(dx) \,\lambda_1(dz).$$

From Lemma 1 we conclude that

$$\forall \sigma > 0 : N(0, \sigma^2) * \mu_1 = N(0, \sigma^2) * \mu_2.$$

Then, by Lemma 2 and Corollary III.3.1,  $\mu_1 = \mu_2$ .

**Example 3.** For independent random variables  $X_1$  and  $X_2$  with  $X_j \sim \pi(\lambda_j)$  we have  $X_1 + X_2 \sim \pi(\lambda_1 + \lambda_2)$ .

Proof: Theorem 2 and Example 1.(i) yield

$$\varphi_{X_1+X_2}(y) = \exp(\lambda_1 \cdot (\exp(\imath y) - 1)) \cdot \exp(\lambda_2 \cdot (\exp(\imath y) - 1))$$
$$= \exp((\lambda_1 + \lambda_2) \cdot (\exp(\imath y) - 1)).$$

Use Theorem 3.

**Lemma 3.** For every  $\varepsilon > 0$  and every probability measure  $\mu \in \mathfrak{M}(\mathbb{R})$ ,

$$\mu(\{x \in \mathbb{R} : |x| \ge 1/\varepsilon\}) \le 7/\varepsilon \cdot \int_0^\varepsilon (1 - \Re \widehat{\mu}(y)) \, dy.$$

*Proof.* Clearly

$$\Re\widehat{\mu}(y) = \int_{\mathbb{R}} \cos(xy) \,\mu(dx).$$

Hence, with the convention  $\sin(0)/0 = 1$ ,

$$1/\varepsilon \cdot \int_0^\varepsilon (1 - \Re\widehat{\mu}(y)) \, dy = 1/\varepsilon \cdot \int_{[0,\varepsilon]} \int_{\mathbb{R}} (1 - \cos(xy)) \, \mu(dx) \, \lambda_1(dy)$$
$$= \int_{\mathbb{R}} \left( 1/\varepsilon \cdot \int_0^\varepsilon (1 - \cos(xy)) \, dy \right) \, \mu(dx)$$
$$= \int_{\mathbb{R}} (1 - \sin(\varepsilon x)/(\varepsilon x)) \, \mu(dx)$$
$$\ge \inf_{|z| \ge 1} (1 - \sin(z)/z) \cdot \mu(\{x \in \mathbb{R} : |\varepsilon x| \ge 1\})$$

Finally,

$$\inf_{|z| \ge 1} (1 - \sin(z)/z) \ge 1/7$$

## Theorem 4 (Continuity Theorem, Lévy).

(i) Let  $\mu, \mu_n \in \mathfrak{M}(\mathbb{R}^k)$  for  $n \in \mathbb{N}$ . Then

$$\mu_n \xrightarrow{w} \mu \qquad \Rightarrow \qquad \forall y \in \mathbb{R}^k : \lim_{n \to \infty} \widehat{\mu_n}(y) = \widehat{\mu}(y).$$

(ii) Let  $\mu_n \in \mathfrak{M}(\mathbb{R}^k)$  for  $n \in \mathbb{N}$ , and let  $\varphi : \mathbb{R}^k \to \mathbb{C}$  be continuous at 0 with  $\varphi(0) = 1$ . Then

$$\forall y \in \mathbb{R}^k : \lim_{n \to \infty} \widehat{\mu_n}(y) = \varphi(y) \qquad \Rightarrow \qquad \exists \mu \in \mathfrak{M}(\mathbb{R}^k) : \widehat{\mu} = \varphi \land \mu_n \stackrel{\mathrm{w}}{\longrightarrow} \mu.$$

*Proof.* Ad (i): Note that  $x \mapsto \exp(i\langle x, y \rangle)$  is bounded and continuous on  $\mathbb{R}^k$ . Ad (ii): See Bauer (1996, Thm. 23.8) or Billingsley (1979, Sec. 29) for the case k > 1. Here: the case k = 1.

We first show that

$$\{\mu_n : n \in \mathbb{N}\} \text{ is tight.}$$
(1)

By Lemma 3

$$\mu_n(\{x \in \mathbb{R} : |x| \ge 1/\varepsilon\}) \le 7/\varepsilon \cdot c_n(\varepsilon)$$

with

$$c_n(\varepsilon) = \int_0^{\varepsilon} (1 - \Re \widehat{\mu_n}(y)) \, dy.$$

Since  $|\Re \widehat{\mu_n}| \leq |\widehat{\mu_n}| \leq 1$ , we have by dominated convergence

$$\lim_{n \to \infty} c_n(\varepsilon) = c(\varepsilon)$$

with

$$c(\varepsilon) = \int_0^{\varepsilon} (1 - \Re \varphi(y)) \, dy$$
.

Now we exploit the assumptions on  $\varphi$ ; given  $\delta > 0$  take  $\varepsilon > 0$  such that

 $7/\varepsilon \cdot c(\varepsilon) \leq \delta/2.$ 

Furthermore, take  $n_0 \in \mathbb{N}$  such that, for every  $n \geq n_0$ ,

$$|c_n(\varepsilon) - c(\varepsilon)| \le \varepsilon/7 \cdot \delta/2.$$

Hence, for  $n \ge n_0$ ,

$$\mu_n(\{x \in \mathbb{R} : |x| \ge 1/\varepsilon\}) \le \delta,$$

and hereby we get (1). Thus, by Prohorov's Theorem,  $\{\mu_n : n \in \mathbb{N}\}$  is relatively compact; on the other hand, for any point of accumulation  $\mu$  of the sequence  $\mu_n$ , Part (i) implies that  $\varphi = \hat{\mu}$ . Hence, there is exactly one point  $\mu$  of accumulation, and for this  $\mu$  we have  $\hat{\mu} = \varphi$ . Finally, Remark 4.4 reveals that  $\mu_n \xrightarrow{w} \mu$ .

**Corollary 2.** Weak convergence in  $\mathfrak{M}(\mathbb{R}^k)$  is equivalent to pointwise convergence of Fourier transforms.