

Chapter III

Basic Concepts of Probability Theory

Context for probability theoretical concepts: a probability space $(\Omega, \mathfrak{A}, P)$.

Terminology: $A \in \mathfrak{A}$ event, $P(A)$ probability of the event $A \in \mathfrak{A}$.

1 Random Variables and Distributions

Given: a probability space $(\Omega, \mathfrak{A}, P)$ and a measurable space (Ω', \mathfrak{A}') .

Definition 1. $X : \Omega \rightarrow \Omega'$ random element if X is \mathfrak{A} - \mathfrak{A}' -measurable. Particular cases:

- (i) X (real) random variable if $(\Omega', \mathfrak{A}') = (\mathbb{R}, \mathfrak{B})$,
- (ii) X numerical random variable if $(\Omega', \mathfrak{A}') = (\overline{\mathbb{R}}, \overline{\mathfrak{B}})$,
- (iii) X k -dimensional (real) random vector if $(\Omega', \mathfrak{A}') = (\mathbb{R}^k, \mathfrak{B}_k)$,
- (iv) X k -dimensional numerical random vector if $(\Omega', \mathfrak{A}') = (\overline{\mathbb{R}}^k, \overline{\mathfrak{B}}_k)$.

Definition 2.

- (i) *Distribution (probability law)* of a random element $X : \Omega \rightarrow \Omega'$ (with respect to P)

$$\mathfrak{L}(X) := P_X(A') := P(\{X^{-1}(A')\}), \quad A' \in \mathfrak{A}'$$

Notation: Q Prob. measure on (Ω', \mathfrak{A}') , then $X \sim Q$ iff $P_X = Q$.

- (ii) Given: probability spaces $(\Omega_1, \mathfrak{A}_1, P_1)$, $(\Omega_2, \mathfrak{A}_2, P_2)$ and random elements

$$X_1 : \Omega_1 \rightarrow \Omega', \quad X_2 : \Omega_2 \rightarrow \Omega'.$$

X_1 and X_2 are *identically distributed* if

$$(P_1)_{X_1} = (P_2)_{X_2}.$$

Remark 1.

- (i) For random elements $X, Y : \Omega \rightarrow \Omega'$

$$X = Y \text{ } P\text{-a.s.} \quad \Rightarrow \quad P_X = P_Y,$$

but the converse is not true in general. For instance, let P be the uniform distribution on $\Omega = \{0, 1\}$ and define $X(\omega) = \omega$ and $Y(\omega) = 1 - \omega$.

- (ii) For every probability measure Q on (Ω', \mathfrak{A}') there exists a probability space $(\Omega, \mathfrak{A}, P)$ and a random element $X : \Omega \rightarrow \Omega'$ such that $X \sim Q$: Choose $(\Omega, \mathfrak{A}, P) = (\Omega', \mathfrak{A}', Q)$ and $X = \text{id}_\Omega$.
- (iii) A major part of probability theory deals with properties of random elements that can be formulated in terms of their distributions.

Example 1.

- (i) *Discrete distributions*, specified by a countable set $\emptyset \neq D \subset \Omega'$ and a mapping $p : D \rightarrow \mathbb{R}$ such that

$$\forall r \in D : p(r) \geq 0 \quad \wedge \quad \sum_{r \in D} p(r) = 1,$$

namely,

$$P_X = \sum_{r \in D} p(r) \cdot \varepsilon_r.$$

Thus, if $\{r\} \in \mathfrak{A}'$ for every $r \in D$,

$$P(\{X = r\}) = p(r).$$

If $|D| < \infty$ then $p(r) = \frac{1}{|D|}$ yields the *uniform distribution on D* .

For $(\Omega', \mathfrak{A}') = (\mathbb{R}, \mathfrak{B})$

$$B(n, p) = \sum_{k=0}^n \binom{n}{k} \cdot p^k (1-p)^{n-k} \cdot \varepsilon_k$$

is the *binomial distribution* with parameters $n \in \mathbb{N}$ and $p \in [0, 1]$. In particular, for $n = 1$ we get the *Bernoulli distribution*

$$B(1, p) = (1-p) \cdot \varepsilon_0 + p \cdot \varepsilon_1.$$

Further examples include the *geometric distribution* with parameter $p \in]0, 1]$,

$$G(p) = \sum_{k=1}^{\infty} p \cdot (1-p)^{k-1} \cdot \varepsilon_k,$$

and the *Poisson distribution* with parameter $\lambda > 0$,

$$\pi(\lambda) = \sum_{k=0}^{\infty} \exp(-\lambda) \cdot \frac{\lambda^k}{k!} \cdot \varepsilon_k.$$

(ii) *Distributions on $(\mathbb{R}^k, \mathfrak{B}_k)$ that are absolutely continuous w.r.t. λ_k , namely, due to the Radon-Nikodym-Theorem*

$$P_X = f \cdot \lambda_k,$$

where

$$f \in \overline{\mathfrak{F}}_+(\mathbb{R}^k, \mathfrak{B}_k) \quad \wedge \quad \int f d\lambda_k = 1.$$

Thus

$$P(\{X \in A'\}) = \int_{A'} f d\lambda_k$$

for every $A' \in \mathfrak{B}_k$.

We present some examples in the case $k = 1$. The *normal distribution*

$$N(\mu, \sigma^2) = f \cdot \lambda_1,$$

with parameters $\mu \in \mathbb{R}$ and σ^2 , where $\sigma > 0$, is obtained by

$$f(x) = \frac{1}{\sqrt{2\pi\sigma^2}} \cdot \exp\left(-\frac{1}{2} \frac{(x - \mu)^2}{\sigma^2}\right), \quad x \in \mathbb{R}.$$

The *exponential distribution* with parameter $\lambda > 0$ is obtained by

$$f(x) = \begin{cases} 0 & \text{if } x < 0 \\ \lambda \cdot \exp(-\lambda x) & \text{if } x \geq 0. \end{cases}$$

The *uniform distribution* on $D \in \mathfrak{B}$ with $\lambda_1(D) \in]0, \infty[$ is obtained by

$$f = \frac{1}{\lambda_1(D)} \cdot 1_D.$$

(iii) Distributions on product spaces can be constructed by means of the results from Section II.8.

Remark 2. Define $\infty^r = \infty$ for $r > 0$. For $1 \leq p < q < \infty$ and $X \in \overline{\mathfrak{F}}(\Omega, \mathfrak{A})$

$$\int |X|^p dP \leq \left(\int |X|^q dP \right)^{p/q},$$

due to Hölder's inequality.

Notation:

$$\mathfrak{L} = \mathfrak{L}(\Omega, \mathfrak{A}, P) = \left\{ X \in \mathfrak{F}(\Omega, \mathfrak{A}) : \int |X| dP < \infty \right\}$$

is the class of P -integrable random variables, and analogously

$$\overline{\mathfrak{L}} = \overline{\mathfrak{L}}(\Omega, \mathfrak{A}, P) = \left\{ X \in \overline{\mathfrak{F}}(\Omega, \mathfrak{A}) : \int |X| dP < \infty \right\}$$

is the class of P -integrable numerical random variables. We consider P_X as a distribution on $(\mathbb{R}, \mathfrak{B})$ if $P(\{X \in \mathbb{R}\}) = 1$ for a numerical random variable X , and we consider \mathfrak{L} as a subspace of $\overline{\mathfrak{L}}$.

Definition 3. For $X \in \overline{\mathfrak{L}}$

$$E(X) = \int X dP$$

is the *expectation* of X . For $X \in \overline{\mathfrak{J}}(\Omega, \mathfrak{A})$ such that $X^2 \in \overline{\mathfrak{L}}$

$$\text{Var}(X) = \int (X - E(X))^2 dP$$

and $\sqrt{\text{Var}(X)}$ are the *variance* and the *standard deviation* of X , respectively.

Remark 3. Theorem II.9.1 implies

$$\int_{\Omega} |X|^p dP < \infty \quad \Leftrightarrow \quad \int_{\mathbb{R}} |x|^p P_X(dx) < \infty$$

for $X \in \overline{\mathfrak{J}}(\Omega, \mathfrak{A})$, in which case, for $p = 1$

$$E(X) = \int_{\mathbb{R}} x P_X(dx),$$

and for $p = 2$

$$\text{Var}(X) = \int_{\mathbb{R}} (x - E(X))^2 P_X(dx).$$

Thus $E(X)$ and $\text{Var}(X)$ depend only on P_X .

Example 2.

$X \sim B(n, p)$	$E(X) = n \cdot p$	$\text{Var}(X) = n \cdot p \cdot (1 - p)$
$X \sim G(p)$	$E(X) = \frac{1}{p}$	$\text{Var}(X) = \frac{1 - p}{p^2}$
$X \sim \pi(\lambda)$	$E(X) = \lambda$	$\text{Var}(X) = \lambda,$

see Introduction to Statistics.

X is *Cauchy distributed* with parameter $\alpha > 0$ if $X \sim f \cdot \lambda_1$ where

$$f(x) = \frac{\alpha}{\pi(\alpha^2 + x^2)}, \quad x \in \mathbb{R}.$$

Since $\int_0^t \frac{x}{1+x^2} dx = \frac{1}{2} \log(1 + t^2)$ neither $E(X^+) < \infty$ nor $E(X^-) < \infty$, and therefore $X \notin \overline{\mathfrak{L}}$.

If $X \sim N(\mu, \sigma^2)$ then

$$E(X) = \mu \quad \wedge \quad \text{Var}(X) = \sigma^2,$$

see Introduction to Statistics.

If X is exponentially distributed with parameter $\lambda > 0$ then

$$E(X) = \frac{1}{\lambda} \quad \wedge \quad \text{Var}(X) = \frac{1}{\lambda^2}.$$

Definition 4. Let $X = (X_1, \dots, X_k)$ be a random vector. Then

$$F_X : \mathbb{R}^k \rightarrow [0, 1]$$

$$(x_1, \dots, x_k) \mapsto P_X \left(\prod_{i=1}^k]-\infty, x_i] \right) = P \left(\bigcap_{i=1}^k \{X_i \leq x_i\} \right)$$

is called the *distribution function* of X .

Theorem 1. Given: probability spaces $(\Omega_1, \mathfrak{A}_1, P_1)$, $(\Omega_2, \mathfrak{A}_2, P_2)$ and random vectors

$$X^1 : \Omega_1 \rightarrow \mathbb{R}^k, \quad X^2 : \Omega_2 \rightarrow \mathbb{R}^k.$$

Then

$$(P_1)_{X^1} = (P_2)_{X^2} \quad \Leftrightarrow \quad F_{X^1} = F_{X^2}.$$

Proof. ‘ \Rightarrow ’ holds trivially. ‘ \Leftarrow ’: $(P_1)_{X^1}$ and $(P_2)_{X^2}(A)$ coincide by assumption on the \cap -stable class

$$\mathfrak{C} = \left\{ \prod_{i=1}^k]-\infty, x_i] : x_1, \dots, x_k \in \mathbb{R} \right\};$$

hence by Theorem II.4.4 they coincide on $\sigma(\mathfrak{C}) = \mathfrak{B}_k$ (see Remark II.1.6). \square

For notational convenience, we consider the case $k = 1$ in the sequel.

Theorem 2.

- (i) F_X is non-decreasing,
- (ii) F_X is right-continuous,
- (iii) $\lim_{x \rightarrow -\infty} F_X(x) = 0$ and $\lim_{x \rightarrow \infty} F_X(x) = 1$,
- (iv) F_X is continuous at x iff $P(\{X = x\}) = 0$.

Proof. Übung 4.1 a). \square

Theorem 3. For every function F that satisfies (i)–(iii) from Theorem 2,

$$\exists_1 Q \text{ probability measure on } \mathfrak{B} : \forall x \in \mathbb{R} : Q(]-\infty, x]) = F(x).$$

Proof. Analogously to the construction of the Lebesgue measure; see Übung 4.1.b). \square